



MTA SZTAKI

Hungarian Academy of Sciences
Institute for Computer Science and Control

Undermodelling Detection with Sign-Perturbed Sums

Algo Carè^{1,2} Marco Campi³ Balázs Csáji² Erik Weyer⁴

¹Centrum Wiskunde & Informatica (CWI), Amsterdam, Netherlands

²Institute for Computer Science and Control (SZTAKI), Hungarian Academy of Sciences (MTA), Hungary

³Department of Information Engineering (DII), University of Brescia, Italy

⁴Department of Electrical and Electronic Engineering, University of Melbourne, Australia

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Motivations

- **SPS** (Sign-Perturbed Sums) builds **confidence regions** around the **LS** (least squares) estimate of **linear regression** problems.
- Only **mild statistical assumptions** are needed, e.g., symmetry.
- Not needed: stationarity, moments, particular distributions.
- SPS has many nice properties (as we will see later), most importantly its confidence regions are **exact**.
- Regarding the models, the assumption of SPS is that the **true system** generating the observations **is in the model class**.
- However, if the model class is wrong, SPS cannot detect it.
- Here, we suggest an extension of SPS, **UD-SPS**, that still builds **exact** confidence sets, if the model is correct, but can also **detect**, in the long run, if the system is **undermodelled**.

Linear Regression

Consider a standard **linear regression** problem:

Linear Regression

$$y_t \triangleq \varphi_t^T \theta^* + w_t$$

where

y_t — **output** (for time $t = 1, \dots, n$)

φ_t — **regressor** (exogenous, d dimensional)

w_t — **noise** (independent, symmetric)

θ^* — **true parameter** (deterministic, d dimensional)

$\Phi_n = [\varphi_1, \dots, \varphi_n]^T$ — skinny and full rank

Least Squares

Given: a sample, \mathcal{Z} , of size n of outputs $\{y_t\}$ and regressors $\{\varphi_t\}$
A classical approach is to minimize the **least squares** criterion

$$\mathcal{V}(\theta | \mathcal{Z}) \triangleq \frac{1}{2} \sum_{t=1}^n (y_t - \varphi_t^T \theta)^2.$$

The **least squares estimate** (LSE) can be found by solving

Normal Equation

$$\nabla_{\theta} \mathcal{V}(\hat{\theta}_n | \mathcal{Z}) = \sum_{t=1}^n \varphi_t (y_t - \varphi_t^T \hat{\theta}_n) = 0$$

Confidence Ellipsoids

LSE is **asymptotically normal** (under some technical conditions)

$$\sqrt{n}(\hat{\theta}_n - \theta^*) \xrightarrow{d} \mathcal{N}(0, \sigma^2 R^{-1}) \text{ as } n \rightarrow \infty,$$

where R is the limit of $R_n = \frac{1}{n} \sum_{t=1}^n \varphi_t \varphi_t^T$ as $n \rightarrow \infty$ (if exists).

Confidence Ellipsoid

$$\tilde{\Theta}_{n,\mu} \triangleq \left\{ \theta \in \mathbb{R}^d : (\theta - \hat{\theta}_n)^T R_n (\theta - \hat{\theta}_n) \leq \frac{\mu \hat{\sigma}_n^2}{n} \right\}$$

where $\mathbb{P}(\theta^* \in \tilde{\Theta}_{n,\mu}) \approx F_{\chi^2(d)}(\mu)$, where $F_{\chi^2(d)}$ is the CDF of $\chi^2(d)$,
 $\hat{\sigma}_n^2 \triangleq \frac{1}{n-d} \sum_{t=1}^n (y_t - \varphi_t^T \hat{\theta}_n)^2$, is an estimate of σ^2 .

Reference and Sign-Perturbed Sums

Let us introduce a **reference sum** and $m - 1$ **sign-perturbed sums**.

Reference Sum

$$S_0(\theta) \triangleq R_n^{-\frac{1}{2}} \sum_{t=1}^n \varphi_t (y_t - \varphi_t^T \theta)$$

Sign-Perturbed Sums

$$S_i(\theta) \triangleq R_n^{-\frac{1}{2}} \sum_{t=1}^n \varphi_t \alpha_{i,t} (y_t - \varphi_t^T \theta)$$

for $i = 1, \dots, m - 1$, where $\alpha_{i,t}$ ($t = 1, \dots, n$) are i.i.d. **random signs**, that is $\alpha_{i,t} = \pm 1$ with probability $1/2$ each (Rademacher).

Intuitive Idea: Distributional Invariance

Recall: $\{w_t\}$ are independent and each w_t is **symmetric** about zero.
Observe that, if $\theta = \theta^*$, we have $(i = 1, \dots, m - 1)$

Distributional Invariance

$$S_0(\theta^*) = R_n^{-\frac{1}{2}} \sum_{t=1}^n \varphi_t w_t$$

$$S_i(\theta^*) = R_n^{-\frac{1}{2}} \sum_{t=1}^n \varphi_t \alpha_{i,t} w_t$$

Consider the **ordering** $\|S_{(0)}(\theta^*)\|^2 \prec \dots \prec \|S_{(m-1)}(\theta^*)\|^2$

Note: relation “ \prec ” is the canonical “ $<$ ” with random tie-breaking

All orderings are equally probable! (they are **conditionally** i.i.d.)

Intuitive Idea: Reference Dominance

What if $\theta \neq \theta^*$?

In fact, the reference paraboloid $\|S_0(\theta)\|^2$ increases faster than $\{\|S_i(\theta)\|^2\}$, thus will eventually **dominate** the ordering.

Intuitively, for “**large enough**” $\|\tilde{\theta}\|$, where $\tilde{\theta} \triangleq \theta^* - \theta$

Eventual Dominance of the Reference Paraboloid

$$\left\| \sum_{t=1}^n \varphi_t \varphi_t^T \tilde{\theta} + \sum_{t=1}^n \varphi_t w_t \right\|_{R_n^{-1}}^2 > \left\| \sum_{t=1}^n \pm \varphi_t \varphi_t^T \tilde{\theta} + \sum_{t=1}^n \pm \varphi_t w_t \right\|_{R_n^{-1}}^2$$

with “**high probability**” (for simplicity \pm is used instead of $\{\alpha_{i,t}\}$).

Non-Asymptotic Confidence Regions

The **rank** of $\|S_0(\theta)\|^2$ in the ordering of $\{\|S_i(\theta)\|^2\}$ w.r.t. \prec is

$$\mathcal{R}(\theta) = 1 + \sum_{i=1}^{m-1} \mathbb{I}(\|S_i(\theta)\|^2 \prec \|S_0(\theta)\|^2),$$

where $\mathbb{I}(\cdot)$ is an indicator function.

Sign-Perturbed Sums (SPS) Confidence Regions

$$\hat{\Theta}_n \triangleq \left\{ \theta \in \mathbb{R}^d : \mathcal{R}(\theta) \leq m - q \right\}$$

where $m > q > 0$ are **user-chosen** integers (design parameters).

Exact Confidence

(A1) $\{w_t\}$ is a sequence of **independent** random variables.

Each w_t has a **symmetric** probability distribution about zero.

(A2) The outer product of regressors is **invertible**, $\det(R_n) \neq 0$.

Exact Confidence of SPS

$$\mathbb{P}(\theta^* \in \hat{\Theta}_n) = 1 - \frac{q}{m}$$

for finite samples. Parameters m and q are under our control.

Note that $\|S_0(\hat{\theta}_n)\|^2 = 0$, thus $\hat{\theta}_n \in \hat{\Theta}_n$, assuming it is non-empty.

Star Convexity

Set $\mathcal{X} \subseteq \mathbb{R}^d$ is **star convex** if there is a **star center** $c \in \mathbb{R}^d$ with

$$\forall x \in \mathcal{X}, \forall \beta \in [0, 1] : \beta x + (1 - \beta) c \in \mathcal{X}.$$

Star Convexity of SPS

$\hat{\Theta}_n$ is star convex with the LSE, $\hat{\theta}_n$, as a star center

Hint $\hat{\Theta}_n$ is the union and intersection of ellipsoids containing LSE.

Strong Consistency

- (A1) **independence, symmetricity**: $\{w_t\}$ are independent, symmetric
- (A2) **invertibility**: $R_n \triangleq \frac{1}{n} \sum_{t=1}^n \varphi_t \varphi_t^T$ is invertible
- (A3) **regressor growth rate**: $\sum_{t=1}^{\infty} \|\varphi_t\|^4 / t^2 < \infty$
- (A4) **noise moment growth rate**: $\sum_{t=1}^{\infty} (\mathbb{E}[w_t^2])^2 / t^2 < \infty$
- (A5) **Cesàro summability**: $\lim_{n \rightarrow \infty} R_n = R$, which is positive definite

Strong Consistency of SPS

$$\mathbb{P} \left(\bigcup_{k=1}^{\infty} \bigcap_{n=k}^{\infty} \left\{ \hat{\Theta}_n \subseteq B_{\varepsilon}(\theta^*) \right\} \right) = 1,$$

where $B_{\varepsilon}(\theta^*) \triangleq \{ \theta \in \mathbb{R}^d : \|\theta - \theta^*\| \leq \varepsilon \}$ is a norm ball.

Ellipsoidal Outer Approximation

The reference paraboloid can be rewritten as

$$\|S_0(\theta)\|^2 = (\theta - \hat{\theta}_n)^T R_n (\theta - \hat{\theta}_n).$$

From which an **alternative** description of the confidence region is

$$\hat{\Theta}_n \subseteq \left\{ \theta \in \mathbb{R}^d : (\theta - \hat{\theta}_n)^T R_n (\theta - \hat{\theta}_n) \leq r(\theta) \right\},$$

where $r(\theta)$ is the q th largest value of $\{\|S_i(\theta)\|^2\}_{i \neq 0}$.

Ellipsoidal Outer Approximation

$$\hat{\Theta}_n \subseteq \left\{ \theta \in \mathbb{R}^d : (\theta - \hat{\theta}_n)^T R_n (\theta - \hat{\theta}_n) \leq r^* \right\}$$

Where r^* can be efficiently computed by a **semi-definite** program.

Undermodelling

Assume we are given a (finite) sample of input and output data, $\{u_t\}$, $\{y_t\}$, which we model with an **FIR** system

$$\hat{y}_t(\theta) \triangleq \varphi_t^T \theta + w_t,$$

where $\varphi_t \triangleq [u_{t-1}, \dots, u_{t-d}]^T$

The true data generation system

$$y_t = \varphi_t^T \theta^* + e_t + n_t,$$

where e_t is an extra component that can depend on **all** past inputs $u_{t-d-1}, u_{t-d-2}, \dots$ and on **all** past noises n_{t-1}, n_{t-2}, \dots

If $\{e_t\}$ are nonzero, then the SPS confidence regions will still (almost surely) shrink, but around a **wrong** parameter value.

SPS with Undermodelling Detection

UD-SPS is obtained from SPS by replacing $\{S_i(\theta)\}$ with

$$Q_0(\theta) \triangleq \begin{bmatrix} R_n & B_n \\ B_n^\top & D_n \end{bmatrix}^{-\frac{1}{2}} \frac{1}{n} \sum_{t=1}^n \begin{bmatrix} \varphi_t \\ \psi_t \end{bmatrix} (y_t - \varphi_t^\top \theta),$$

$$Q_i(\theta) \triangleq \begin{bmatrix} R_n & B_n \\ B_n^\top & D_n \end{bmatrix}^{-\frac{1}{2}} \frac{1}{n} \sum_{t=1}^n \alpha_{i,t} \begin{bmatrix} \varphi_t \\ \psi_t \end{bmatrix} (y_t - \varphi_t^\top \theta),$$

where ψ_t is a vector that includes **s extra input values** preceding the \hat{n}_b that are included in φ_t , $\psi_t \triangleq [u_{t-d-1}, \dots, u_{t-d-s}]^\top$, and

$$B_n \triangleq \frac{1}{n} \sum_{t=1}^n \varphi_t \psi_t^\top, \quad D_n \triangleq \frac{1}{n} \sum_{t=1}^n \psi_t \psi_t^\top.$$

The Connection of UD-SPS and SPS

The connection of UD-SPS and SPS can be stated as

Reducing UD-SPS to SPS

The **UD-SPS region**, $\hat{\Theta}_n^o$, for estimating $\theta^* \in \mathbb{R}^d$ can be interpreted as the **restriction** to a d -dimensional space of a **standard SPS region**, $\hat{\Theta}'_n$, that lives in the domain $\{\theta' \in \mathbb{R}^{d+s}\}$.

\mathbb{R}^{d+s} is the d -dimensional identification space **augmented** with s extra components: $\hat{\Theta}_n^o$ can be identified with the first d components of the set $\hat{\Theta}'_n \cap (\mathbb{R}^d \times \{0\}^s)$.

UD-SPS with Correct System Specifications

Theorem (Exact Confidence of UD-SPS)

If the FIR system is *correctly specified*, then

$$\mathbb{P}\{\theta^* \in \hat{\Theta}_n^o\} = 1 - q/m.$$

Theorem (Strong Consistency of UD-SPS)

If the FIR system is *correctly specified*, then (under some technical conditions) for all $\varepsilon > 0$, we have that

$$\mathbb{P}\left[\bigcup_{\bar{n}=1}^{\infty} \bigcap_{n=\bar{n}}^{\infty} \{\hat{\Theta}_n^o \subseteq B_\varepsilon(\theta^*)\}\right] = 1,$$

where $B_\varepsilon(\theta^*)$ denotes an ε -ball centred around θ^* .

UD-SPS in the Presence of Undermodelling

Theorem (Undermodelling Detection)

Assume that the system is *undermodelled*, that is $\{e_t\}$ are nonzero (and some technical conditions hold). With the notations

$$\bar{R}' \triangleq \lim_{n \rightarrow \infty} \begin{bmatrix} R_n & B_n \\ B_n^\top & D_n \end{bmatrix}, \quad \bar{E}' \triangleq \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{t=1}^n \begin{bmatrix} \varphi_t \\ \psi_t \end{bmatrix} \mathbb{E}[e_t],$$

if the following *detectability* condition holds

$$\bar{R}'^{-1} \bar{E}' \notin \mathbb{R}^{\hat{n}_b} \times \{0\}^s,$$

then

$$\mathbb{P} \left[\bigcup_{\bar{n}=1}^{\infty} \bigcap_{n=\bar{n}}^{\infty} \{ \hat{\Theta}_n^o = \emptyset \} \right] = 1.$$

Numerical Experiments

Consider the following **ARX**(1,1) data generating system

$$y_t = a^* y_{t-1} + b^* u_{t-1} + n_t,$$

with zero initial conditions, where $a^* = 0.5$ or 0.15 or 0 (see later), $b^* = 1$, $\{n_t\}$ are i.i.d. Laplacian with mean 0 and variance 0.1 .

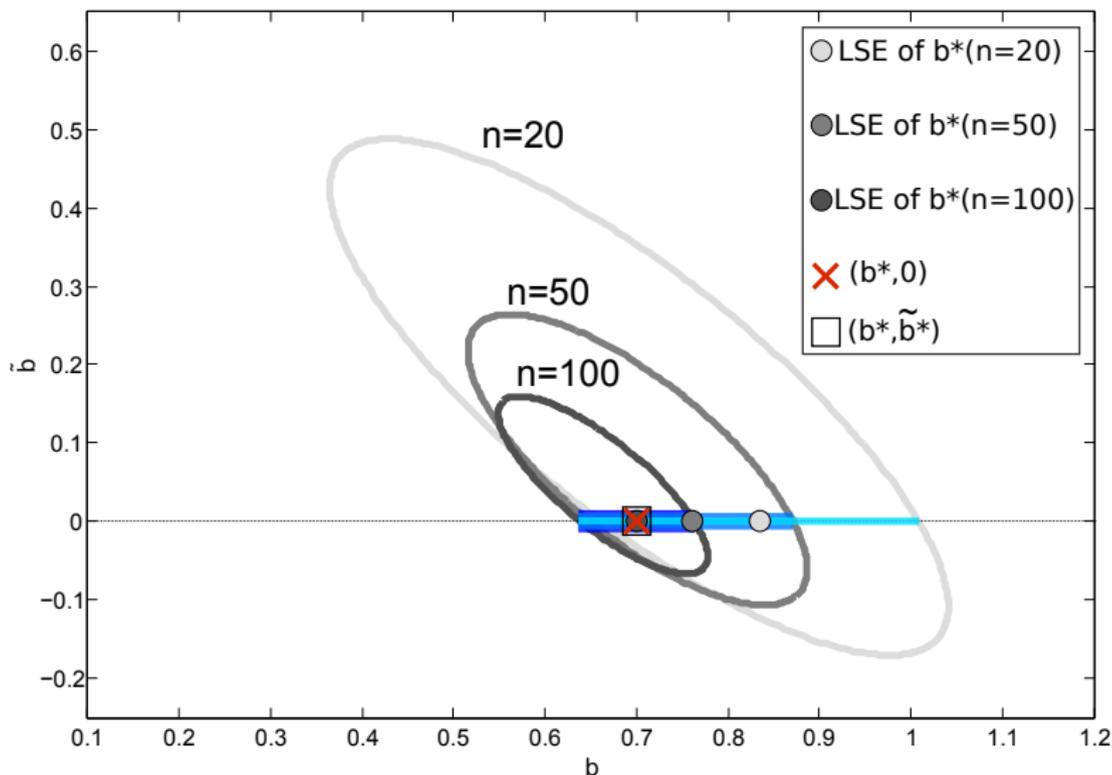
The **input signal** is generated as $u_t = 0.75u_{t-1} + v_t$, where $\{v_t\}$ are i.i.d. standard normal random variables.

The user-chosen predictor is an **FIR**(1) model

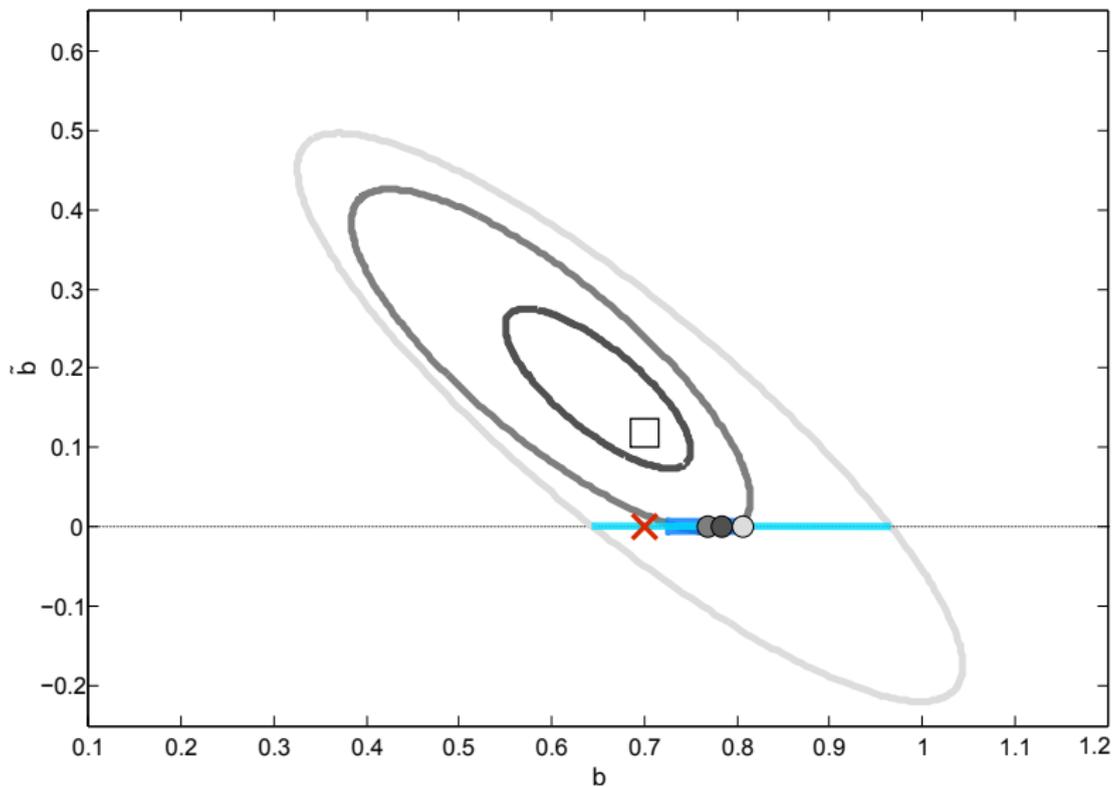
$$\hat{y}_t(\theta) = \varphi_t^\top \theta = b u_{t-1},$$

that is, the **autoregressive part is missing**, $\theta = [b]$ is the model parameter, and $\varphi_t = [u_{t-1}]$ is the regressor at time t .

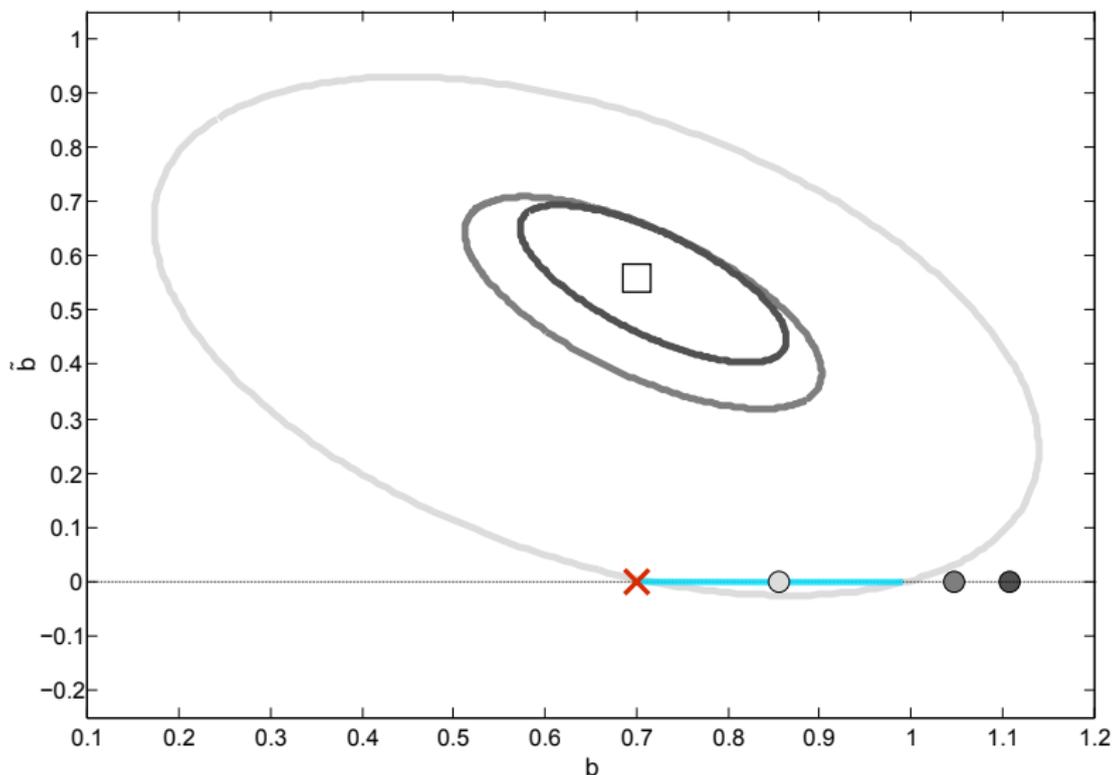
95% UD-SPS Confidence Intervals, $a^* = 0$



95% UD-SPS Confidence Intervals, $a^* = 0.15$



95% UD-SPS Confidence Intervals, $a^* = 0.5$



Summary and Conclusions

- **SPS** (Sign-Perturbed Sums) is a powerful finite sample system identification method that builds **exact**, **star convex**, **strongly consistent** confidence regions for linear regression problems.
- SPS also has efficient **ellipsoidal outer-approximations**.
- However, SPS **cannot detect** if the model class is wrong.
- Here, we suggested an extension of SPS, called **UD-SPS**, that still guarantees **exact** and **strongly consistent** confidence regions if the model order is correctly specified.
- Furthermore, it can **detect**, in the long run, if the model is undermodelled (detection = empty confidence region).
- There is a strong **connection** between SPS and UD-SPS.
- The theoretical results were also confirmed by numerical **experiments**: FIR models of ARX systems were studied.

Thank you for your attention!

✉ balazs.csaji@sztaki.mta.hu